

A note on randomly colored matchings in random bipartite graphs

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Abstract

We are given a bipartite graph that contains at least one perfect matching and where each edge is colored from a set $Q = \{c_1, c_2, \dots, c_q\}$. Let $Q_i = \{e \in E(G) : c(e) = c_i\}$, where $c(e)$ denotes the color of e . The perfect matching color profile $mcp(G)$ is defined to be the set of vectors $(m_1, m_2, \dots, m_q) \in [n]^q$ such that there exists a perfect matching M such that $|M \cap Q_i| = m_i$. We give bounds on the matching color profile for a randomly colored random bipartite graph.

1 Introduction

We consider the following problem: we are given a random bipartite graph G in which each edge is given a random color from a set $Q = \{c_1, c_2, \dots, c_q\}$. An edge e is colored $c(e) = c_i$ with probability α_i where $\alpha_i > 0$ is a constant. Let $Q_i = \{e \in E(G) : c(e) = c_i\}$, where $c(e)$ denotes the color of e . The perfect matching color profile $mcp(G)$ is defined to be the set of vectors $(m_1, m_2, \dots, m_q) \in [n]^q$ such that there exists a perfect matching M such that $|M \cap Q_i| = m_i$. We give bounds on the matching color profile for a randomly colored random bipartite graph.

Randomly colored random graphs have been studied recently in the context of (i) rainbow matchings and Hamilton cycles, see for example [2], [3], [7], [11]; (ii) rainbow connection see for example [5],

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[9], [10], [13], [12]; (iii) pattern colored Hamilton cycles, see for example [1], [6]. This paper can be considered to be a contribution in the same genre. One can imagine a possible interest in the color profile via the following scenario: suppose that A is a set of tools and B is a set of jobs where edge $\{a, b\}$ indicates that b can be completed using a . If colors represent people, then one might be interested in equitably distributing jobs. I.e. determining whether $(n/q, n/q, \dots, n/q) \in mcp(G)$. In any case, we find the problem interesting.

We will consider G to be the random bipartite graph $G_{n,n,p}$ where $p = \frac{\log n + \omega}{n}$, $\omega = \omega(n) \rightarrow \infty$ where $\omega = o(\log n)$. Erdős and Rényi [4] proved that G has a perfect matching w.h.p. We will prove the following theorem: let $\alpha_1, \alpha_2, \dots, \alpha_q, \beta$ be positive constants such that $\alpha_1 + \alpha_2 + \dots + \alpha_q = 1$ and $\beta < 1/q$. Let

$$\alpha_{\min} = \min \{\alpha_i : i \in [q]\}.$$

Theorem 1. *Let G be the random bipartite graph $G_{n,n,p}$ where $p = \frac{\log n + \omega}{n}$, $\omega = \omega(n) \rightarrow \infty$ where $\omega = o(\log n)$. Suppose that the edges of G are independently colored with colors from $C = \{c_1, c_2, \dots, c_q\}$ where $\mathbb{P}(c(e) = c_i) = \alpha_i$ for $e \in E(G), i \in [q]$. Let m_1, m_2, \dots, m_q satisfy: (i) $m_1 + \dots + m_q = n$ and (ii) $m_i \geq \beta n, i \in [q]$. Then w.h.p., there exists a perfect matching M in which exactly m_i edges are colored with $c_i, i = 1, 2, \dots, q$.*

It is clear that w.h.p. $(n, 0, \dots, 0) \notin mcp(G)$. This is because the bipartite graph induced by edges of color c_1 is distributed as $G_{n,n,\alpha_1 p}$ and this contains isolated vertices w.h.p. On the other hand, if $p \geq \frac{q(\log n + \omega)}{\alpha_{\min} n}$ then w.h.p. $mcp(G) = [n]^q$. To see this, suppose that $m_1 \leq m_2 \leq \dots \leq m_q \leq n$. Suppose we have found a matching that uses m_i edges of color c_i for $i \geq 0$. Let $n' = n - m_1 - \dots - m_i$. Then the random bipartite graph induced by vertices not in M and having edges of color c_i has density at least $\frac{q\alpha_i n'}{\alpha_{\min} n} \cdot \frac{\log n + \omega}{n'} \geq \frac{\log n' + \omega/2}{n'}$ and so has a perfect matching w.h.p.

Open Question: What is the threshold for $mcp(G) = [0, n]^q$?

2 Structural Lemma

Suppose that the bipartition of $V(G)$ is denoted A, B . For sets $S \subseteq A, T \subseteq B$ we let $e_i(S, T)$ denote the number of $S : T$ edges of color c_i . We say that vertex u is c_i -adjacent to vertex v if the edge $\{u, v\}$ exists and has color c_i .

Lemma 2. *Let $p = \frac{\log n + \omega}{n}$, $\omega = \omega(n) \rightarrow \infty$ where $\omega = o(\log n)$. Then w.h.p.*

- (a) $S \subseteq A, T \subseteq B$ and $\gamma_a \log n \leq |S| \leq n_0 = \gamma_a n / \log n$ and $|T| \leq \alpha_i \eta |S| \log n$ where $\gamma_a = \eta / (20\alpha_i)$ implies that $e_i(S : T) \leq 2\alpha_i \eta |S| \log n$ for $i = 1, 2, \dots, q$.
- (b) There do not exist sets $X \subseteq S \subseteq A, T \subseteq B$ and $i \in [q]$ such that $|S|, |T| \geq \beta n$ and $|X| = \gamma_b |S| / \log n$, $\gamma_b = 10 \log(e/\beta) / \alpha_i$ and such that each $x \in X$ is c_i -adjacent to fewer than $\alpha_i \beta \log n / 10$ vertices in T .

- (c) There do not exist sets $X \subseteq S \subseteq A, T \subseteq B$ and $i \in [q]$ such that $|S|, |T| \geq \beta n$ and $|X| = |S|/\log n$ and a set $Z \subseteq T, |Z| = \gamma_b n / \log n$ such that each $x \in X$ is c_i -adjacent to $k = \frac{10 \log n}{\log \log n}$ vertices in Z .
- (d) There do not exist sets $S \subseteq A, T \subseteq B$ and $i \in [q]$ such that $|S|, |T| \geq \beta n$ such that there are more than $\gamma_d n / \log n, \gamma_d = \frac{4}{\alpha_i} \log\left(\frac{e}{\beta}\right)$ vertices in T that not c_i -adjacent to a vertex in S .
- (e) Fix $\gamma, \delta > 0$ constants. Then w.h.p. there do not exist sets S, T with $|S| = |T| = \gamma n / \log n$ such that $e_i(S, T) \geq \delta |S| \log n / \log \log n$.
- (f) There do not exist sets $S \subseteq A, T \subseteq B$ and $i \in [q]$ such that $|S|, |T| \geq \beta n / 10$ such that $e_i(S, T) = 0$.

Proof

(a) The probability that the condition is violated can be bounded by

$$\begin{aligned}
& \sum_{s=\gamma_a \log n}^{n_0} \sum_{t=1}^{\alpha_i \eta s \log n} \binom{n}{s} \binom{n}{t} \binom{st}{2\alpha_i \eta s \log n} (\alpha_i p)^{2\alpha_i \eta s \log n} \\
& \leq \sum_{s=\gamma_a \log n}^{n_0} \sum_{t=1}^{\alpha_i \eta s \log n} \left(\frac{ne}{s}\right)^s \left(\frac{ne}{t}\right)^t \left(\frac{est\alpha_i p}{2\alpha_i \eta s \log n}\right)^{2\alpha_i \eta s \log n} \\
& \leq \sum_{s=\gamma_a \log n}^{n_0} \sum_{t=1}^{\alpha_i \eta s \log n} \left(\frac{ne}{s}\right)^s \left(\frac{ne}{\alpha_i \eta s \log n}\right)^{\alpha_i \eta s \log n} \left(\frac{etp}{2\eta \log n}\right)^{2\alpha_i \eta s \log n} \\
& \leq \sum_{s=\gamma_a \log n}^{n_0} \sum_{t=1}^{\alpha_i \eta s \log n} \left(\left(\frac{ne}{s}\right)^{1/2\alpha_i \eta \log n} \left(\frac{ne}{\alpha_i \eta s \log n}\right)^{1/2} \cdot \frac{e^{1+o(1)} \alpha_i s \log n}{2n}\right)^{2\alpha_i \eta s \log n} \\
& \leq \sum_{s=\gamma_a \log n}^{n_0} \sum_{t=1}^{\alpha_i \eta s \log n} \left(\frac{s \log n}{n}\right)^{\alpha_i \eta s \log n - s} (\log n)^s \left(\frac{e^{3/2+o(1)} \alpha_i^{1/2}}{2\eta^{1/2}}\right)^{2\alpha_i \eta s \log n} = o(1).
\end{aligned}$$

(b) The probability that the condition is violated can be bounded by

$$\binom{n}{\beta n}^2 \binom{\beta n}{\gamma_b n / \log n} (e^{-\alpha_i \beta / 4})^{\gamma_b n} \leq \left(\left(\frac{e}{\beta}\right)^{(2\beta+o(1))} e^{-\alpha_i \beta \gamma_b / 4}\right)^n = o(1).$$

The factor $e^{-\alpha_i \beta \gamma_b / 4}$ comes from applying a Chernoff bound.

(c) We can assume w.l.o.g. that $|S| = |T| = \beta n$. The probability that the condition is violated can be bounded by

$$\begin{aligned}
& \binom{n}{\beta n}^2 \binom{\beta n}{n / \log n} \binom{\beta n}{\gamma_b n / \log n} \left(\binom{\gamma_b n / \log n}{k} (\alpha_i p)^k\right)^{n / \log n} \\
& \leq \left(\frac{e}{\beta}\right)^{(2\beta+o(1))n} \left(\frac{e \gamma_b \alpha_i}{k}\right)^{kn / \log n} = o(1).
\end{aligned}$$

(d) The probability that the condition is violated can be bounded by

$$\binom{n}{\beta n}^2 \binom{\beta n}{\gamma_d n / \log n} (1 - \alpha_i p)^{\beta n \gamma_d n / \log n} \leq \left(\left(\frac{e}{\beta} \right)^{2+o(1)} e^{-\alpha_i \gamma_d} \right)^{\beta n} = o(1).$$

(e) The probability that the condition is violated can be bounded by

$$\binom{n}{\gamma n / \log n}^2 \left(\frac{\gamma^2 n^2 / (\log n)^2}{\delta n / \log \log n} \right) p^{\delta n / \log \log n} \leq \left(\frac{e \log n}{\gamma} \right)^{2\gamma n / \log n} \left(\frac{\gamma^2 e \log \log n}{\delta \log n} \right)^{\delta n / \log \log n} = o(1).$$

(f) The probability that the condition is violated can be bounded by

$$2^{2n} (1 - p)^{\beta^2 n^2 / 100} = o(1).$$

□

3 Proof of Theorem 1

Proof Assume from now on that the high probability conditions of Lemma 2 are in force. Let M be a perfect matching and let $\mu_i = |M \cap Q_i|$ for $i \in [q]$. Suppose that $\mu_1 > m_1 \geq \beta n$ and $\beta n \leq \mu_2 < m_2$. We show that we can find another matching M' such that $|M' \cap Q_1| = \mu_1 - 1$ and $|M' \cap Q_2| = \mu_2 + 1$. We do this by finding an alternating cycle with edge sequence $C = (e_1, f_1, \dots, e_\ell, f_\ell)$ and vertex sequence $(x_1 \in A, y_1 \in B, x_2, \dots, x_\ell, y_\ell, x_1)$ such that (i) $e_i = \{x_i, y_i\} \in M$, (ii) $f_i = \{y_i, x_{i+1}\} \notin M, i \in [\ell]$, (iii) $e_1 \in Q_1$ and (iv) $E(C) \setminus \{e_1\} \subseteq Q_2$. Repeating this for pairs of colors, one over-subscribed and one under-subscribed we eventually achieve our goal. It is sufficient to consider this case, seeing as we can always w.h.p. find a matching that has been randomly colored with $\approx \alpha_i n$ edges of color $c_i, i = 1, 2, \dots, q$.

Next let $A_i = V(M \cap Q_i) \cap A$ and $B_i = V(M \cap Q_i) \cap B$ for $i \in [q]$ and for $S \subseteq A$ let $N_i(S) = \{b \in B : \exists a \in S \text{ s.t. } \{a, b\} \in Q_i\}$ and $N_i(a) = N_i(\{a\})$. Then let

$$D'_0 = \left\{ a \in A_2 : |N_2(a) \cap B_2| \geq \frac{\alpha_2 \beta \log n}{10} \right\}.$$

$$D_0 = \left\{ a \in A_1 : |N_2(a) \cap M(A_2 \setminus D'_0)| \leq k_0 = \frac{10 \log n}{\log \log n} \right\}.$$

It follows from Lemma 2(b) that

$$|M(A_2 \setminus D'_0)| \leq \frac{\gamma_b n}{\log n}.$$

It then follows from Lemma 2(c) that if $W_0 = A_1 \setminus D_0$ then

$$|W_0| \leq \frac{n}{\log n}. \tag{1}$$

We now define a sequence of sets W_0, W_1, \dots where W_{j+1} is obtained from W_j by adding a vertex of $A_2 \setminus W_j$ for which $|N_2(a) \cap M(W_j)| \geq k_0$. Now consider $S = W_t, T = M(W_t)$ for some $t \geq 1$. Then we have

$$|S| = |T| \leq t + \frac{n}{\log n} \text{ and } e_2(S, T) \geq tk_0.$$

Given Lemma 2(e) with $\delta = 5, \gamma = 2$, we see that this sequence stops with $t = t^* \leq 4n/\log n$. So we now let $R_0 = A_2 \setminus W_{t^*}$. We note that

$$\begin{aligned} |R_0| &\geq \beta n - \frac{5n}{\log n} \\ a \in R_0 \text{ implies } |N_2(a) \cap M(R_0)| &\geq \frac{\alpha_2 \beta \log n}{10} - k_0. \end{aligned} \tag{2}$$

We now fix some $a_0 \in R_0$ and define a sequence of sets $X_0, Y_0, X_1, Y_1, \dots$ where $X_j \subseteq R_0$ and $Y_j \subseteq B_2$. We let $X_0 = \{a_0\}$ and then having defined $X_i, i \geq 0$ we let

$$Y_i = N_2(X_i) \text{ and } X_{i+1} = \left(M^{-1}(Y_i) \setminus \bigcup_{j \leq i} X_j \right) \cap R_0.$$

We claim that for $i \geq 0$,

$$|X_i| \leq \frac{n}{200 \log n} \text{ implies that } |X_{i+1}| \geq \frac{\alpha_2 \beta \log n}{25} |X_i|. \tag{3}$$

We verify (3) below. Assuming its truth, there exists a smallest k such that

$$|X_k| \geq \frac{\alpha_2 \beta n}{5000}. \tag{4}$$

Starting with $\hat{Y}_0 = \{b_0\}$ where $b_0 = M(a_0) \in \hat{R}_0$, we can similarly construct a sequence of sets $\hat{Y}_1, \hat{X}_1, \dots$ where $\hat{X}_j \subseteq M^{-1}(\hat{R}_0)$ and $\hat{Y}_j \subseteq \hat{R}_0$. Here \hat{R}_0 is the equivalently defined set to R_0 in B_2 . We can assume that $b_0 \in \hat{R}_0$, because of the sizes of the sets R_0, \hat{R}_0 . More precisely, by (1), there will be $o(n)$ choices for a_0 for which $b_0 \notin \hat{R}_0$. Having defined \hat{Y}_i we let

$$\hat{X}_i = N_2(\hat{Y}_i) \text{ and } \hat{Y}_{i+1} = \left(M(\hat{X}_i) \setminus \bigcup_{j \leq i} \hat{Y}_j \right) \cap \hat{R}_0.$$

and then let $\hat{Y}_{i+1} = M(\hat{X}_i)$. The equivalent of (3) will be

$$|\hat{Y}_i| \leq \frac{n}{200 \log n} \text{ implies that } |\hat{Y}_{i+1}| \geq \frac{\alpha_2 \beta \log n}{25} |\hat{Y}_i|. \tag{5}$$

Assuming its truth, there exists ℓ such that

$$|\hat{Y}_\ell| \geq \frac{\alpha_2 \beta n}{5000}. \tag{6}$$

It follows from Lemma 2(f) that at least 9/10 of the vertices of A_1 have a c_2 -neighbor in \hat{R}_0 and at least 9/10 of the vertices of B_1 have a c_2 -neighbor in R_0 . We deduce from this that there is a pair $x_0 \in A_1, y_0 = M(x_0) \in B_1$ such that $N_2(x_0) \cap \hat{R}_0 \neq \emptyset$ and $N_2(y_0) \cap R_0 \neq \emptyset$. This defines an alternating cycle $x_0, u_0, P_1, b_0, a_0, P_2, v_0, y_0, x_0$. Here u_0 is a c_2 -neighbor of x_0 in \hat{R}_0 and P_1 is (the reversal of) a path from u_0 to b_0 and P_2 is the path from a_0 to $v_0 \in X_k, v_0 \in N_2(y_0)$. This completes the proof of Theorem 1.

Verification of (3), (5): We have by the assumption $a_0 \in R_0$ that

$$|X_1| = |Y_1| \geq \frac{\alpha_2 \beta \log n}{10} - o(\log n)$$

Now suppose that $1 \leq |X_i| \leq n/(200 \log n)$. Then, by (2),

$$e_2(X_i : (N_2(X_i) \setminus M(A_2 \setminus R_0))) \geq \frac{(\alpha_2 \beta \log n)|X_i|}{10 + o(1)}.$$

Applying Lemma 2(a) we see that

$$|N_2(X_i) \setminus M(A_2 \setminus R_0)| \geq \frac{(\alpha_2 \beta \log n)|X_i|}{20 + o(1)}. \quad (7)$$

Because the sets X_1, X_2, \dots expand rapidly, the total size of $\bigcup_{j \leq i} X_j$ is small compared with the R.H.S of (7) and (3) follows. The argument for (5) is similar. \square

4 Concluding Remarks

We have established that w.h.p. $mcp(G)$ is almost all of $[0, n]^q$ and posed the question of finding the exact threshold for $mcp(G) = [0, n]^q$. It seems technically feasible to extend our results to randomly colored $G_{n,p}$. We leave this for future research. It would be of some interest to analyse other spanning subgraphs from this point of view e.g. Hamilton cycles.

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