# A geometric preferential attachment model of networks 

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#### Abstract

We study a random graph $G_{n}$ that combines certain aspects of geometric random graphs and preferential attachment graphs. The vertices of $G_{n}$ are $n$ sequentially generated points $x_{1}, x_{2}, \ldots, x_{n}$ chosen uniformly at random from the unit sphere in $\boldsymbol{R}^{3}$. After generating $x_{t}$, we randomly connect it to $m$ points from those points in $x_{1}, x_{2}, \ldots, x_{t-1}$ which are within distance $r$. Neighbors are chosen with probability proportional to their current degree and a parameter $\alpha$ biasses the choice towards self loops. We show that if $m$ is sufficiently large, if $r \geq \ln n / n^{1 / 2-\beta}$ for some constant $\beta$, and if $\alpha>2$, then whp at time $n$ the number of vertices of degree $k$ follows a power law with exponent $\alpha+1$. Unlike the preferential attachment graph, this geometric preferential attachment graph has small separators, similar to experimental observations of [8]. We further show that if $m \geq K \ln n, K$ sufficiently large, then $G_{n}$ is connected and has diameter $O(\ln n / r)$ whp.


## 1 Introduction

Recently there has been much interest in understanding the properties of real-world large-scale networks such as the structure of the Internet and the World Wide Web. For a general introduction to this topic, see Bollobás and Riordan [9], Hayes [23], Watts [34], or Aiello, Chung and Lu [3]. One approach is to model these networks by random graphs. Experimental studies by Albert, Barabási, and Jeong [4], Broder et al [13], and Faloutsos, Faloutsos, and Faloutsos [21] have demonstrated that in the World Wide Web/Internet the proportion of vertices of a given degree follows an approximate inverse power law i.e. the proportion of vertices of degree $k$ is approximately $C k^{-\alpha}$ for some constants $C, \alpha$. The classical models of random graphs introduced by Erdős and Renyi [19] do not have power law degree sequences, so they are not suitable for modeling these networks. This has driven the development of various alternative models for random graphs.

[^0]One approach is to generate graphs with a prescribed degree sequence (or prescribed expected degree sequence). This is proposed as a model for the web graph by Aiello, Chung, and Lu in [1]. Mihail and Papadimitriou also use this model [29] in their study of large eigenvalues, as do Chung, Lu , and Vu in [15].
An alternative approach, which we will follow in this paper, is to sample graphs via some generative procedure which yields a power law distribution. There is a long history of such models, outlined in the survey by Mitzenmacher [31]. We will use an extension of the preferential attachment model to generate our random graph. The preferential attachment model has been the subject of recently revived interest. It dates back to Yule [35] and Simon [33]. It was proposed as a random graph model for the web by Barabási and Albert [5], and their description was elaborated by Bollobás and Riordan [10] who showed that at time $n$, whp the diameter of a graph constructed in this way is asymptotic to $\frac{\ln n}{\ln \ln n}$. Subsequently, Bollobás, Riordan, Spencer and Tusnády [12] proved that the degree sequence of such graphs does follow a power law distribution.
The random graph defined in the previous paragraph has good expansion properties. For example, Mihail, Papadimitriou and Saberi [30] showed that whp the preferential attachment model has conductance bounded below by a constant. On the other hand, Blandford, Blelloch and Kash [8] found that some WWW related graphs have smaller separators than what would be expected in random graphs with the same average degree. The aim of this paper is to describe a random graph model which has both a power-law degree distribution and which has small separators.
We study here the following process which generates a sequence of graphs $G_{t}, t=$ $1,2, \ldots, n$. The graph $G_{t}=\left(V_{t}, E_{t}\right)$ has $t$ vertices and $m t$ edges. Here $V_{t}$ is a subset of $S$, the surface of the sphere in $\boldsymbol{R}^{3}$ of radius $\frac{1}{2 \sqrt{\pi}}$ (so that $\operatorname{area}(S)=1$ ).
For $u \in S$ and $r>0$ we let $B_{r}(u)$ denote the spherical cap of radius $r$ around $u$ in $S$. More precisely, $B_{r}(u)=\{x \in S:\|x-u\| \leq r\}$.

### 1.1 The random process

The parameters of the process are $m>0$ the number of edges added in every step and $\alpha \geq 0$ a measure of the bias towards self loops.
Notice that there exists a constant $c_{0}$ such that for any $u \in S$, we have

$$
A_{r}=\operatorname{Area}\left(B_{r}(u)\right) \sim c_{0} r^{2}
$$

- Time step 0: To initialize the process, we start with $G_{0}$ being the Empty Graph.
- Time step $t+1$ : We choose vertex $x_{t+1}$ uniformly at random in $S$ and add it to $G_{t}$. Let $V_{t}\left(x_{t}\right)=V_{t} \cap B_{r}\left(x_{t+1}\right)$ and let $D_{t}\left(x_{t}\right)=\sum_{v \in V_{t}\left(x_{t}\right)} \operatorname{deg}_{t}(v)$. We add $m$ random edges $\left(x_{t+1}, y_{i}\right), i=1,2, \ldots, m$ incident with $x_{t+1}$. Here, each $y_{i}$ is chosen independently from $V_{t}\left(x_{t}\right) \cup\left\{x_{t+1}\right\}$ (parallel edges and loops are permitted), such that for each $i=1, \ldots, m$, for all $v \in V_{t}\left(x_{t+1}\right)$,

$$
\operatorname{Pr}\left(y_{i}=v\right)=\frac{\operatorname{deg}_{t}(v)}{\max \left(D_{t}\left(x_{t+1}\right), \alpha m A_{r} t\right)}
$$

and

$$
\operatorname{Pr}\left(y_{i}=x_{t+1}\right)=1-\frac{D_{t}\left(x_{t+1}\right)}{\max \left(D_{t}\left(x_{t+1}\right), \alpha m A_{r} t\right)}
$$

(When $t=0$ we have $\operatorname{Pr}\left(y_{i}=x_{1}\right)=1$.)
Let $d_{k}(t)$ denote the number of vertices of degree $k$ at time $t$ and let $\bar{d}_{k}(t)$ denote the expectation of $d_{k}(t)$.
We will prove the following:

## Theorem 1

(a) If $0<\beta<1 / 2$ and $\alpha>2$ are constants and $r \sim n^{\beta-1 / 2} \ln n$ and $m$ is a sufficiently large constant then there exist constants $c, \gamma, \epsilon>0$ such that for all $k=k(n) \geq m$,

$$
\begin{equation*}
\bar{d}_{k}(n)=C_{k} \frac{n}{k^{1+\alpha}}+O\left(n^{1-\gamma}\right) \tag{1}
\end{equation*}
$$

where $C_{k}=C_{k}(m, \alpha)$ tends to a constant $C_{\infty}(m, \alpha)$ as $k \rightarrow \infty$.
Furthermore, for $n$ sufficiently large, the random variable $d_{k}(n)$ satisfies the following concentration inequality:

$$
\begin{equation*}
\operatorname{Pr}\left(\left|d_{k}(n)-\bar{d}_{k}(n)\right| \geq n^{1-\gamma}\right) \leq e^{-n^{\epsilon}} \tag{2}
\end{equation*}
$$

(b) If $\alpha \geq 0$ and $r=o(1)$ then $\mathbf{w h p} V_{n}$ can be partitioned into $T, \bar{T}$ such that $|T|,|\bar{T}| \sim$ $n / 2$, and there are at most $4 \sqrt{\pi} r n m$ edges between $T$ and $\bar{T}$.
(c) If If $\alpha \geq 0$ and $r \geq n^{-1 / 2} \ln n$ and $m \geq K \ln n$ and $K$ is sufficiently large then whp $G_{n}$ is connected.
(d) If If $\alpha \geq 0$ and $r \geq n^{-1 / 2} \ln n$ and $m \geq K \ln n$ and $K$ is sufficiently large then whp $G_{n}$ has diameter $O(\ln n / r)$.

We note that geometric models of trees with power laws have been considered in [20], [6] and [7]. We also note that Gómez-Gardeñes and Moreno [22] have empirically analyzed a one dimensional version of our model when $\alpha=0$ and their experiments suggest that this yields a power-law exponent of 3 .

### 1.2 Open Questions

In an earlier version of the paper there was no $\alpha$ and we have failed to produce a proof of Theorem 1 (a) when $\alpha \leq 2$. This remains a challenge for us at the present moment. We do not think that the $\ln n$ factors are necessary in parts (c),(d).

### 1.3 Some definitions

Given $U \subseteq S$ and $u \in S$, we define

$$
V_{t}(U)=V_{t} \cap U \quad \text { and } \quad V_{t}(u)=V_{t}\left(B_{r}(u)\right)
$$

and

$$
D_{t}(U)=\sum_{v \in V_{t}(U)} \operatorname{deg}_{t}(v) \quad \text { and } \quad D_{t}(u)=D_{t}\left(B_{r}(u)\right)
$$

Given $v \in V_{t}$, we also define

$$
\begin{equation*}
\operatorname{deg}_{t}^{-}(v)=\operatorname{deg}_{t}(v)-m \tag{3}
\end{equation*}
$$

Notice that $\operatorname{deg}_{t}^{-}(v)$ is the number of edges of $G_{t}$ that are incident to $v$ and were added by vertices that chose $v$ as a neighbor, including loops at $v$.
Given $U \subseteq S$, let $D_{t}^{-}(U)=\sum_{v \in V_{t}(U)} \operatorname{deg}_{t}^{-}(v)$. We also define $D_{t}^{-}(u)=D_{t}^{-}\left(B_{r}(u)\right)$.
Notice that $D_{t}(U)=m\left|V_{t}(U)\right|+D_{t}^{-}(U)$.
We localize some of our notation: given $U \subseteq S$ and $u \in S$ we define $d_{k}(t, U)$ to be the number of vertices of degree $k$ at time $t$ in $U$ and $d_{k}(t, u)=d_{k}\left(t, B_{r}(u)\right)$.

## 2 Outline of the paper

In Section 3 we show that there are small separators. This is easy, since any give great circle can whp be used to define a small separator.
We prove a likely power law for the degree sequence in Section 4. We follow a standard practise and prove a recurrence for the expected number of vertices of degree $k$ at time step $t$. Unfortunatley, this involves the estimation of the expectation of the reciprocal of a random variable and to handle this, we show that this random variable is concentrated. This is quite technical and is done in Section 4.3.
Section 5 proves connectivity when $m$ grows logarithmically with $n$. The idea is to show that whp the sub-graph $G_{n}(B)$ induced by a ball $B$ of radius $r / 2$, center $u \in$ $S$, is connected. This is done by constructing a connected subgraph of $G_{n}(B)$ via a coupling argument. We then show that the union of the $G_{n}(B)$ 's for $u=x_{1}, x_{2}, \ldots, x_{n}$ is connected and has small diameter.

## 3 Small separators

Theorem 1(b) is the easiest part to prove. We use the geometry of the instance to obtain a sparse cut. Consider partitioning the vertices using a great circle of $S$. This will divide $V$ into sets $T$ and $\bar{T}$ which each contain about $n / 2$ vertices. More precisely, we have

$$
\operatorname{Pr}[|T|<(1-\epsilon) n / 2]=\operatorname{Pr}[|\bar{T}|<(1-\epsilon) n / 2] \leq e^{-\epsilon^{2} n / 4}
$$

Edges only appear between vertices within distance $r$, so only vertices appearing in the strip within distance $r$ of the great circle can appear in the cut. Since $r=o(1)$, this
strip has area less than $3 r \sqrt{\pi}$, and, letting $U$ denote the vertices appearing in this strip, we have

$$
\operatorname{Pr}[|U| \geq 4 \sqrt{\pi} r n] \leq e^{-\sqrt{\pi} r n / 9}
$$

Even if every one of the vertices chooses its $m$ neighbors on the opposite side of the cut, this will yield at most $4 \sqrt{\pi} r n m$ edges whp. So the graph has a cut with $\frac{e(T, \bar{T})}{|T||T|} \leq \frac{17 \sqrt{\pi} r m}{n}$ with probability at least $1-e^{-\Omega(r n)}$.

## 4 Proving a power law

### 4.1 Establishing a recurrence for $\bar{d}_{k}(t)$ : the expected number of vertices of degree $k$ at time $t$

Our approach to proving Theorem 1(a) is to find a recurrence for $\bar{d}_{k}(t)$.
We define $\bar{d}_{m-1}(t)=0$ for all integers $t$ with $t>0$. Let $\eta_{k}\left(G_{t}, x_{t+1}\right)$ denote the (conditional) probability that a parallel edge to a vertex of degree no more than $k$ is created. Then,

$$
\begin{array}{r}
\eta_{k}\left(G_{t}, x_{t+1}\right)=O\left(\sum_{i=m}^{k} \frac{d_{i}\left(t, x_{t+1}\right) i^{2}}{\max \left\{\alpha m A_{r} t, D_{t}\left(x_{t+1}\right)\right\}^{2}}\right) \\
=O\left(\min \left\{\frac{k^{2}}{\max \left\{\alpha m A_{r} t, D_{t}\left(x_{t+1}\right)\right\}}, 1\right\}\right) \tag{4}
\end{array}
$$

Then for $k \geq m$,

$$
\begin{align*}
& \mathbf{E}\left[d_{k}(t+1) \mid G_{t}, x_{t+1}\right]= \\
& \quad d_{k}(t) \\
& \quad+m d_{k-1}\left(t, x_{t+1}\right) \frac{k-1}{\max \left\{\alpha m A_{r} t, D_{t}\left(x_{t+1}\right)\right\}}-m d_{k}\left(t, x_{t+1}\right) \frac{k}{\max \left\{\alpha m A_{r} t, D_{t}\left(x_{t+1}\right)\right\}}  \tag{5}\\
& \quad+\operatorname{Pr}\left[d e g_{t+1}\left(x_{t+1}=k\right) \mid G_{t}, x_{t+1}\right]+O\left(m \eta_{k}\left(G_{t}, x_{t+1}\right)\right) .
\end{align*}
$$

Let $\mathcal{A}_{t}$ be the event

$$
\left\{\left|D_{t}\left(x_{t+1}\right)-2 m A_{r} t\right| \leq C_{1} A_{r} m t^{\gamma} \ln n\right\}
$$

where

$$
\max \{2 / \alpha, 1 / 2,1-2 \beta\}<\gamma<1
$$

and $C_{1}$ is some sufficiently large constant.
Note that if

$$
t \geq(\ln n)^{2 /(1-\gamma)}
$$

then

$$
\mathcal{A}_{t} \text { implies } D_{t}\left(x_{t+1}\right) \leq \alpha m A_{r} t .
$$

Then, because $\mathbf{E}\left[d_{k}\left(t, x_{t+1}\right)\right] \leq k^{-1} \mathbf{E}\left[m\left|V_{t}\left(B_{2 r}\left(x_{t+1}\right)\right)\right|\right] \leq k^{-1} m\left(4 A_{r} t\right)$ and $d_{k}\left(t, x_{t+1}\right) \leq$ $k^{-1} D_{t}\left(x_{t+1}\right)<m t$, we have for $t \geq(\ln n)^{2 /(1-\gamma)}$,

$$
\begin{aligned}
& \mathbf{E}\left[\frac{d_{k}\left(t, x_{t+1}\right)}{\max \left\{\alpha m A_{r} t, D_{t}\left(x_{t+1}\right)\right\}}\right] \\
& =\quad \mathbf{E}\left[\left.\frac{d_{k}\left(t, x_{t+1}\right)}{\max \left\{\alpha m A_{r} t, D_{t}\left(x_{t+1}\right)\right\}} \right\rvert\, \mathcal{A}_{t}\right] \operatorname{Pr}\left[\mathcal{A}_{t}\right]+ \\
& \\
& +\mathbf{E}\left[\left.\frac{d_{k}\left(t, x_{t+1}\right)}{\max \left\{\alpha m A_{r} t, D_{t}\left(x_{t+1}\right)\right\}} \right\rvert\, \neg \mathcal{A}_{t}\right] \operatorname{Pr}\left[\neg \mathcal{A}_{t}\right] \\
& \quad=\frac{\mathbf{E}\left[d_{k}\left(t, x_{t+1}\right) \mid \mathcal{A}_{t}\right]}{\alpha m A_{r} t} \operatorname{Pr}\left[\mathcal{A}_{t}\right]+\mathbf{E}\left[\left.O\left(\frac{d_{k}\left(t, x_{t+1}\right)}{D_{t}\left(x_{t+1}\right)}\right) \right\rvert\, \neg \mathcal{A}_{t}\right] \operatorname{Pr}\left[\neg \mathcal{A}_{t}\right] \\
& \quad=\frac{\mathbf{E}\left[d_{k}\left(t, x_{t+1}\right) \mid \mathcal{A}_{t}\right]}{\alpha m A_{r} t} \operatorname{Pr}\left[\mathcal{A}_{t}\right]+O\left(\frac{\operatorname{Pr}\left[\neg \mathcal{A}_{t}\right]}{k}\right) \\
& \quad=\frac{\mathbf{E}\left[d_{k}\left(t, x_{t+1}\right)\right]}{\alpha m A_{r} t}+\left(O\left(\frac{1}{k}\right)-\frac{\mathbf{E}\left[d_{k}\left(t, x_{t+1}\right) \mid \neg \mathcal{A}_{t}\right]}{\alpha m A_{r} t}\right) \operatorname{Pr}\left[\neg \mathcal{A}_{t}\right] \\
& \quad=\frac{\mathbf{E}\left[d_{k}\left(t, x_{t+1}\right)\right]}{\alpha m A_{r} t}+O\left(\frac{1}{k}+\frac{1}{A_{r}}\right) \mathbf{P r}\left[\neg \mathcal{A}_{t}\right] .
\end{aligned}
$$

In Lemmas 1 and 3 below we prove that

$$
\mathbf{E}\left[d_{k}\left(t, x_{t+1}\right)\right]=m A_{r} \bar{d}_{k}(t)
$$

and that

$$
\begin{equation*}
\operatorname{Pr}\left[\neg \mathcal{A}_{t}\right]=O\left(n^{-2}\right) . \tag{6}
\end{equation*}
$$

Thus, if $t \geq(\ln n)^{2 /(1-\gamma)}$ then

$$
\begin{equation*}
\mathbf{E}\left[\frac{d_{k}\left(t, x_{t+1}\right)}{\max \left\{\alpha m A_{r} t, D_{t}\left(x_{t+1}\right)\right\}}\right]=\frac{\bar{d}_{k}(t)}{\alpha m t}+O\left(\frac{1}{n^{2}}\left(\frac{1}{A_{r}}+\frac{1}{k}\right)\right) . \tag{7}
\end{equation*}
$$

In a similar way

$$
\begin{equation*}
\mathbf{E}\left[\frac{d_{k-1}\left(t, x_{t+1}\right)}{\max \left\{\alpha m A_{r} t, D_{t}\left(x_{t+1}\right)\right\}}\right]=\frac{\bar{d}_{k-1}(t)}{\alpha m t}+O\left(\frac{1}{n^{2}}\left(\frac{1}{A_{r}}+\frac{1}{k}\right)\right) . \tag{8}
\end{equation*}
$$

On the other hand, given $G_{t}, x_{t+1}$, if

$$
p=1-\frac{D_{t}\left(x_{t+1}\right)}{\max \left(D_{t}\left(x_{t+1}\right), \alpha m A_{r} t\right)}
$$

then

$$
\operatorname{Pr}\left[d e g_{t+1}\left(x_{t+1}=k\right) \mid G_{t}, x_{t+1}\right]=\operatorname{Pr}[\operatorname{Bi}(m, p)=k-m]
$$

So, if $t \geq(\ln n)^{2 /(1-\gamma)}$,

$$
\begin{aligned}
& \operatorname{Pr}\left[x_{t+1}=k\right]=\binom{m}{k-m} \mathbf{E}\left[p^{k-m}(1-p)^{2 m-k} \mid \mathcal{A}_{t}\right] \operatorname{Pr}\left[\mathcal{A}_{t}\right]+O\left(\operatorname{Pr}\left[\neg \mathcal{A}_{t}\right]\right) \\
& =\binom{m}{k-m}\left(1-\frac{2}{\alpha}\right)^{k-m}\left(\frac{2}{\alpha}\right)^{2 k-m}\left(1+O\left(t^{\gamma-1} \ln n\right)\right) \operatorname{Pr}\left[\mathcal{A}_{t}\right]+O\left(n^{-2}\right) \\
& =\binom{m}{k-m}\left(1-\frac{2}{\alpha}\right)^{k-m}\left(\frac{2}{\alpha}\right)^{2 k-m}+O\left(t^{\gamma-1} \ln n\right) .
\end{aligned}
$$

Now note that from equations (4) and (6) that if

$$
t \geq t_{0}=n^{(1-2 \beta) / \gamma}
$$

and

$$
k \leq k_{0}(t)=\left(m A_{r} t^{\gamma} \ln n\right)^{1 / 2}
$$

then

$$
\begin{equation*}
\mathbf{E}\left(\eta_{k}\left(G_{t}, x_{t+1}\right)\right)=O\left(t^{\gamma-1} \ln n\right) . \tag{9}
\end{equation*}
$$

Taking expectations on both sides of (5) and using (7,8,9), we see that if $t \geq t_{0}$ and $k \leq k_{0}(t)$ then

$$
\begin{align*}
& \bar{d}_{k}(t+1)=\bar{d}_{k}(t)+\frac{k-1}{\alpha t} \bar{d}_{k-1}(t)-\frac{k}{\alpha t} \bar{d}_{k}(t) \\
&+\binom{m}{k-m}\left(1-\frac{2}{\alpha}\right)^{k-m}\left(\frac{2}{\alpha}\right)^{2 m-k}+O\left(t^{\gamma-1} \ln n\right) \tag{10}
\end{align*}
$$

We consider the recurrence given by $f_{m-1}=0$ and for $k \geq m$,

$$
f_{k}=\frac{k-1}{\alpha} f_{k-1}-\frac{k}{\alpha} f_{k}+\binom{m}{k-m}\left(1-\frac{2}{\alpha}\right)^{k-m}\left(\frac{2}{\alpha}\right)^{2 m-k}
$$

which, for $k>2 m$, has solution

$$
\begin{aligned}
f_{k} & =f_{2 m} \prod_{i=m+1}^{k} \frac{i-1}{i+\alpha} \\
& =\phi_{k}(m, \alpha)\left(\frac{m}{k}\right)^{\alpha+1},
\end{aligned}
$$

and has that $\phi_{k}(m, \alpha)$ tends to a limit $\phi_{\infty}(m, \alpha)$ depending only on $m, \alpha$ as $k \rightarrow \infty$, We can absorb the values $f_{m}, f_{m+1}, \ldots, f_{2 m}$ into this notation.
We finish the proof of (1) by showing that there exists a constant $M>0$ such that

$$
\begin{equation*}
\left|\bar{d}_{k}(t)-f_{k} t\right| \leq M\left(t_{0}+t^{\gamma} \ln n\right) \tag{11}
\end{equation*}
$$

for all $0 \leq t \leq n$ and $m \leq k \leq k_{0}(t)$.
This is trivially true for $t<t_{0}$.
For $k>k_{0}(t)$ this follows from $\bar{d}_{k}(t) \leq 2 m t / k$.
Let $\Theta_{k}(t)=\bar{d}_{k}(t)-f_{k} t$. Then for $t \geq t_{0}$ and $m \leq k \leq k_{0}(t)$,

$$
\begin{equation*}
\Theta_{k}(t+1)=\frac{k-1}{\alpha t} \Theta_{k-1}(t)-\frac{k}{\alpha t} \Theta_{k}(t)+O\left(t^{\gamma-1} \ln n\right) . \tag{12}
\end{equation*}
$$

Let $L$ denote the hidden constant in $O\left(t^{\gamma-1} \ln n\right)$ of (12). Our inductive hypothesis $\mathcal{H}_{t}$ is that

$$
\left|\Theta_{k}(t)\right| \leq M\left(t_{0}+t^{\gamma} \ln n\right)
$$

for every $m \leq k \leq k_{0}(t)$ and $M$ sufficiently large. It is trivially true for $t \leq t_{0}$. So assume that $t \geq t_{0}$. Then, from (12),

$$
\begin{aligned}
\left|\Theta_{k}(t+1)\right| & \leq M\left(t_{0}+t^{\gamma} \ln n\right)+L t^{\gamma-1} \ln n \\
& \leq M\left(t_{0}+(t+1)^{\gamma} \ln n\right) .
\end{aligned}
$$

This verifies $\mathcal{H}_{t+1}$ and completes the proof by induction.

### 4.2 Expected Value of $d_{k}(t, u)$

Lemma 1 Let $u \in S$ and let $k$ and $t$ be positive integers. Then $\mathbf{E}\left[d_{k}(t, u)\right]=A_{r} \bar{d}_{k}(t)$
Proof By symmetry, for any $w \in S, d_{k}(t, u)$ has the same distribution as $d_{k}(t, w)$. Then

$$
\begin{aligned}
\mathbf{E}\left[d_{k}(t, u)\right] & =\int_{S} \mathbf{E}\left[d_{k}(t, u)\right] d w=\int_{S} \mathbf{E}\left[d_{k}(t, w)\right] d w \\
& =\mathbf{E}\left[\int_{S} d_{k}(t, w) d w\right]=\mathbf{E}\left[\int_{S} \sum_{v \in V_{t}} 1_{\operatorname{deg} v=k} 1_{v \in B_{r}(w)} d w\right] \\
& =\mathbf{E}\left[\sum_{v \in V_{t}} 1_{\operatorname{deg} v=k} \int_{S} 1_{w \in B_{r}(v)} d w\right]=\mathbf{E}\left[\sum_{v \in V_{t}} 1_{\operatorname{deg} v=k} A_{r}\right] \\
& =A_{r} \mathbf{E}\left[d_{k}(t)\right]
\end{aligned}
$$

Lemma 2 Let $u \in S$ and $t>0$ then $\mathbf{E}\left[D_{t}(u)\right]=2 A_{r} m t$
Proof

$$
\mathbf{E}\left[D_{t}(u)\right]=\sum_{k>0} \mathbf{E}\left[d_{k}(t, u)\right]=A_{r} \sum_{k>0} \mathbf{E}\left[d_{k}(t)\right]=A_{r} \mathbf{E}\left[\sum_{k>0} d_{k}(t)\right]=2 A_{r} m t
$$

### 4.3 Concentration of $D_{t}(u)$

In this section we prove
Lemma 3 If $t>0$ and $u$ is chosen randomly from $S$ then

$$
\operatorname{Pr}\left[\left|D_{t}(u)-\mathbf{E}\left[D_{t}(u)\right]\right| \geq A_{r} m\left(t^{2 / \alpha}+t^{1 / 2} \ln t\right) \ln n\right]=O\left(n^{-2}\right)
$$

Proof We think of every edge added as two directed edges. We also think of $x_{t}$, the vertex added, as being added with $\left(\alpha m A_{r} t-D_{t}\left(x_{t}\right)\right)^{+}=\max \left\{\alpha m A_{r} t-D_{t}\left(x_{t}\right), 0\right\}$ "phantom" edges pointing to it. Then choosing a vertex is equivalent to choosing one of these directed edges uniformly, and taking the vertex pointed to by this edge as the chosen vertex. So the $i$-th step of the process is defined by a tuple of random variables $T=\left(X, Y_{1}, \ldots, Y_{m}\right) \in S \times E_{i}^{m}$ where $X$ is the location of the new vertex, a randomly chosen point in $S$, and $Y_{j}$ is an edge chosen uniformly at random from among the edges directed into $B_{r}(X)$ in $G_{i-1}$. The process $G_{t}$ is then defined by a sequence $\left\langle T_{1}, \ldots, T_{t}\right\rangle$, where each $T_{i} \in S \times E_{i}^{m}$.
Let $s$ be a sequence $s=\left\langle s_{1}, \ldots, s_{t}\right\rangle$ where $s_{i}=\left(x_{i}, y_{(i-1) m+1}, \ldots, y_{i m}\right)$ with $x_{i} \in S$ and $y_{j} \in E_{\lceil j / m\rceil}$. We say $s$ is acceptable if for every $j, y_{j}$ is an edge entering $B_{r}\left(x_{\lceil t / j\rceil}\right)$. Notice that non-acceptable sequences have probability 0 of being realized. Fix $t>$
0. Fix an acceptable sequence $s=\left\langle s_{1}, \ldots, s_{t}\right\rangle$, and let $A_{\tau}(s)=\left\{z \in S \times E_{\tau}^{m}\right.$ : $\left\langle s_{1}, \ldots, s_{\tau-1}, z\right\rangle$ is acceptable $\}$. For any $\tau$ with $1 \leq \tau \leq t$ and any $z \in A_{\tau}(s)$ let

$$
g_{\tau}(z)=\mathbf{E}\left[D_{t}(u) \mid T_{1}=s_{1}, \ldots, T_{\tau-1}=s_{\tau-1}, T_{\tau}=z\right]
$$

let $r_{\tau}(s)=\sup \left\{\left|g_{\tau}(z)-g_{\tau}(\hat{z})\right|: z, \hat{z} \in A_{\tau}(s)\right\}$ and let $\hat{r}^{2}(s)=\sum_{\tau=1}^{t}\left(\sup _{s} r_{\tau}(s)\right)^{2}$, where the supremum is taken over all acceptable sequences.
From the Azuma-Hoeffding inequality (see for example [2]) we know that for all $\lambda>0$,

$$
\begin{equation*}
\operatorname{Pr}\left[\left|D_{t}(u)-\mathbf{E}\left[D_{t}(u)\right]\right| \geq \lambda\right]<2 e^{-\lambda^{2} / 2 \hat{r}^{2}} \tag{13}
\end{equation*}
$$

Fix $\tau$, with $1 \leq \tau \leq t$. Our goal now is to bound $r_{\tau}(s)$ for any acceptable sequence $s$. Fix $z, \hat{z} \in A_{\tau}(s)$. We define $\Omega\left(G_{t}, \hat{G}_{t}\right)$, a coupling between $G_{t}=G_{t}\left(s_{1}, \ldots, s_{\tau-1}, z\right)$ and $\hat{G}_{t}=G_{t}\left(s_{1}, \ldots, s_{\tau-1}, \hat{z}\right)$

- Step $\tau$ : Start with the graph $G_{\tau}\left(s_{1}, \ldots, s_{\tau-1}, z\right)$ and $\hat{G}_{\tau}\left(s_{1}, \ldots, s_{\tau-1}, \hat{z}\right)$ respectively.
- Step $\sigma(\sigma>\tau)$ : Choose the same point $x_{\sigma} \in S$ in both processes. Let $E_{\sigma}$ (resp. $\left.\hat{E}_{\sigma}\right)$ be the edges pointing to the vertices in $B_{r}\left(x_{\sigma}\right)$ in $G_{\sigma-1}$ (resp. $\hat{G}_{\sigma-1}$ ) plus the $\left(\alpha m A_{r} \sigma-D_{\sigma}\left(x_{\sigma}\right)\right)^{+}\left(\right.$resp. $\left.\left(\alpha m A_{r} \sigma-\hat{D}_{\sigma}\left(x_{\sigma}\right)\right)^{+}\right)$phantom edges pointing to $x_{\sigma}$. Let $C_{\sigma}=E_{\sigma} \cap \hat{E}_{\sigma}, R_{\sigma}=E_{\sigma} \backslash \hat{E}_{\sigma}$, and $L_{\sigma}=\hat{E}_{\sigma} \backslash E_{\sigma}$
Notice that $\left|E_{\sigma}\right|,\left|\hat{E}_{\sigma}\right| \geq \alpha m A_{r} \sigma$. Notice also that if $D_{\sigma}\left(x_{\sigma}\right), D_{\sigma}^{\prime}\left(x_{\sigma}\right) \leq \alpha m A_{r} \sigma$, then $\left|E_{\sigma}\right|=\left|\hat{E}_{\sigma}\right|$ and $\left|R_{\sigma}\right|=\left|L_{\sigma}\right|$. Without loss of generality assume that $\left|E_{\sigma}\right| \leq$ $\left|\hat{E}_{\sigma}\right|$.
Now, define $p=1 /\left|E_{\sigma}\right|$ and $\hat{p}=1 /\left|\hat{E}_{\sigma}\right|$. Construct $G_{\sigma}$ by choosing $m$ edges uniformly at random $e_{1}^{\sigma}, \ldots, e_{m}^{\sigma}$ in $E_{\sigma}$, and then joining $x_{\sigma}$ to their endpoints, $y_{1}^{\sigma}, \ldots, y_{m}^{\sigma}$. For each of the $m$ edges $e_{i}=e_{i}^{\sigma}$, we define $\hat{e}_{i}=\hat{e}_{i}^{\sigma}$ by
- If $e_{i} \in C_{\sigma}$ then, with probability $\hat{p} / p, \hat{e}_{i}=e_{i}$. With probability $1-\hat{p} / p, \hat{e}_{i}$ is chosen from $L_{\sigma}$ uniformly at random.
- If $e_{i} \in R_{\sigma}, \hat{e}_{i} \in L_{\sigma}$ is chosen uniformly at random.

Notice that for every $i=1, \ldots, m$ and every $e \in \hat{E}_{\sigma}, \operatorname{Pr}\left[\hat{e}_{i}=e\right]=\hat{p}$. To finish, in $\hat{G}_{\sigma}$ join $x_{\sigma}$ to the $m$ vertices pointed to by the edges $\hat{e}_{i}$.

Now let

$$
\Delta_{\sigma}=\sum_{\rho=\tau}^{\sigma} \sum_{i=1}^{m} 1_{y_{i}^{\sigma} \neq \hat{y}_{i}^{\sigma}}
$$

and for $u \in S$ let

$$
\Delta_{\sigma}(u)=\sum_{\rho=\tau}^{\sigma} \sum_{i=1}^{m} 1_{\left|\left\{y_{i}^{\sigma}, \hat{y}_{i}^{\sigma}\right\} \cap B_{r}(u)\right|=1}
$$

## Lemma 4

$$
\left|g_{\tau}(z)-g_{\tau}(\hat{z})\right| \leq \mathbf{E}\left[\Delta_{t}(u)\right]
$$

## Proof

$$
\begin{aligned}
\left|g_{\tau}(z)-g_{\tau}(\hat{z})\right| & =\left|\mathbf{E}_{G_{t}}\left[D_{t}(u)\right]-\mathbf{E}_{\hat{G}_{t}}\left[D_{t}(u)\right]\right| \\
& =\left|\mathbf{E}_{\Omega\left(G_{t}, G_{t}^{\prime}\right)}\left[D_{t}(u)-D_{t}^{\prime}(u)\right]\right| \\
& \left.\left.\leq \mathbf{E}_{\Omega\left(G_{t}, G_{t}^{\prime}\right)}\right) \Delta_{t}(u)\right]
\end{aligned}
$$

since only when $\left|\left\{y_{i}^{\sigma}, \hat{y}_{i}^{\sigma}\right\} \cap B_{r}(u)\right|=1$ do we add $\pm 1$ to the difference $D_{\rho}(u)-D_{\rho}^{\prime}(u)$.
Recall that $A_{r}=\operatorname{Area}\left(B_{r}(u)\right) \sim c_{0} n^{2 \beta-1}(\ln n)^{2}$ and we have fixed $\tau$ to be an integer with $1 \leq \tau \leq t$.

Lemma 5 Let $t \geq 1$ and $u \in S$. Then for some constant $C>0$,

$$
\mathbf{E}\left[\Delta_{t}(u)\right] \leq C m A_{r}\left(\frac{t}{\tau}\right)^{2 / \alpha}
$$

Proof Let $\tau<\sigma \leq t$. We start with

$$
\begin{equation*}
\Delta_{\sigma}=\Delta_{\sigma-1}+\sum_{i=1}^{m} 1_{y_{i}^{\sigma} \neq \hat{y}_{i}^{\sigma}} . \tag{14}
\end{equation*}
$$

Now fix $G_{\sigma-1}, \hat{G}_{\sigma-1}$ and $x_{\sigma}$ and $i$. Then taking expectations with respect to our coupling,

$$
\begin{align*}
& \mathbf{E}\left[1_{y_{i}^{\sigma} \neq \hat{y}_{i}^{\sigma}}\right]=\operatorname{Pr}\left(y_{i}^{\sigma} \neq \hat{y}_{i}^{\sigma}\right)=\operatorname{Pr}\left(e_{i}^{\sigma} \neq \hat{e}_{i}^{\sigma}\right)= \\
& 1-\frac{\left|C_{\sigma}\right|}{\left|E_{\sigma}\right|} \frac{\hat{p}}{p}=1-\frac{\left|C_{\sigma}\right|}{\left|\hat{E}_{\sigma}\right|}=\frac{\left|L_{\sigma}\right|}{\left|\hat{E}_{\sigma}\right|}=\frac{\max \left\{\left|L_{\sigma}\right|,\left|R_{\sigma}\right|\right\}}{\max \left\{\left|E_{\sigma}\right|,\left|\hat{E}_{\sigma}\right|\right\}} \leq \frac{\left|L_{\sigma}\right|+\left|R_{\sigma}\right|}{\alpha m A_{r} \sigma} \tag{15}
\end{align*}
$$

Therefore

$$
\begin{equation*}
\mathbf{E}\left[\Delta_{\sigma} \mid G_{\sigma-1}, \hat{G}_{\sigma-1}, x_{\sigma}\right] \leq \Delta_{\sigma-1}+m \frac{\left|L_{\sigma}\right|+\left|R_{\sigma}\right|}{\alpha m A_{r} \sigma} \tag{16}
\end{equation*}
$$

For each $e \in E\left(\hat{G}_{\sigma-1}\right) \backslash E\left(G_{\sigma-1}\right), e \in L_{\sigma}$ implies $x_{\sigma}$ is in the ball of radius $r$ centered at the end point of $e$. Similarly for $e \in R_{\sigma}$. Therefore,

$$
\begin{equation*}
\mathbf{E}\left[\left|L_{\sigma}\right|+\left|R_{\sigma}\right| \mid G_{\sigma-1}, \hat{G}_{\sigma-1}\right] \leq 2 A_{r} \Delta_{\sigma-1} . \tag{17}
\end{equation*}
$$

Then,
$\mathbf{E}\left[\Delta_{\sigma}\right] \leq \mathbf{E}\left[\Delta_{\sigma-1}\right]+m \frac{\mathbf{E}\left[\left|L_{\sigma}\right|+\left|R_{\sigma}\right|\right]}{\alpha m A_{r} \sigma} \leq \mathbf{E}\left[\Delta_{\sigma-1}\right]+\frac{2 \mathbf{E}\left[\Delta_{\sigma-1}\right]}{\alpha \sigma}=\mathbf{E}\left[\Delta_{\sigma-1}\right]\left(1+\frac{2}{\alpha \sigma}\right)$,
so, $\mathbf{E}\left[\Delta_{t}\right] \leq e^{10 / \alpha^{2}}\left(\frac{t}{\tau}\right)^{2 / \alpha} \mathbf{E}\left[\Delta_{\tau}\right]$. Now, $\Delta_{\tau} \leq m$, because the graphs $G_{\tau}$ and $\hat{G}_{\tau}$ differ at most in the last $m$ edges. Therefore $\mathbf{E}\left[\Delta_{t}\right] \leq m e^{10 / \alpha^{2}}\left(\frac{t}{\tau}\right)^{2 / \alpha}$.
Finally, note that if $v$ is a random point in $S$ then $\mathbf{E}\left[\Delta_{t}(v)\right]=A_{r} \mathbf{E}\left[\Delta_{t}\right]$. For this, fix $u$ and let $\phi$ denote a random rotation of $S$. Let $v=\phi(u)$ and then run Process 1 with $\phi\left(G_{\tau}\right), \phi\left(\hat{G}_{\tau}\right)$ and $x_{\sigma}, \sigma>\tau$ and then consider Process 2 starting with $G_{\tau}, \hat{G}_{\tau}$ and
$\phi^{-1}\left(x_{\sigma}\right), \sigma>\tau$. The mapping $\phi^{-1}$ does not disturb the distribution of $x_{\sigma}, \sigma>\tau$ and therefore $\Delta_{t}(u)$ in Process 2 is equal to $\Delta_{t}(v)$ in Process 1.

By applying Lemma 5 , we have that for any acceptable sequence

$$
R^{2}(s)=\sum_{\tau=1}^{t} r_{\tau}(s)^{2} \leq\left(C m A_{r}\right)^{2} t^{4 / \alpha} \sum_{\tau=1}^{t} \tau^{-4 / \alpha}=O\left(A_{r}^{2} m^{2}\left(t \ln t+t^{4 / \alpha}\right)\right)
$$

Therefore, by using Equation (13), we have that there is $C_{1}$ such that

$$
\operatorname{Pr}\left[\left|D_{t}(u)-\mathbf{E}\left[D_{t}(u)\right]\right| \geq C_{1} A_{r} m\left(t^{2 / \alpha}+t^{1 / 2} \ln t\right)(\ln n)^{1 / 2}\right] \leq e^{-2 \ln n}=n^{-2} .
$$

### 4.4 Concentration of $d_{k}(t)$

We follow the proof of Lemma 3, replacing $D_{t}(u)$ by $d_{k}(t)$ and using the same coupling. When we reach Lemma 4 we find that $\left|g_{\tau}(z)-g_{\tau}(\hat{z})\right| \leq 2 \mathbf{E}\left[\widehat{D}_{t}\right]$ (each edge disrepancy can affect two vertices), the rest is the same.
This proves (1) and completes the proof of Theorem 1(a) .

## 5 Connectivity

Here we are going to prove that for $r \geq n^{-1 / 2} \ln n, m>K \ln n$, and $K$ sufficiently large, whp $G_{n}$ is connected and has diameter $O(\ln n / r)$. Notice that $G_{n}$ is a subgraph of the graph $G(n, r)$, the intersection graph of the caps $B_{r}\left(x_{t}\right), t=1,2, \ldots, n$ and therefore it is disconnected for $r=o\left(\left(n^{-1} \ln n\right)^{1 / 2}\right)$ [32]. We denote the diameter of $G$ by $\operatorname{diam}(G)$, and follow the convention of defining $\operatorname{diam}(G)=\infty$, when $G$ is disconnected. In particular, when we say that a graph has finite diameter this implies it is connected.
Let

$$
T=K_{1} \ln n / A_{r}=O(n / \ln n)
$$

where $K_{1}$ is sufficiently large, and $K_{1} \ll K$.
Lemma 6 Let $u \in S$ and let $B=B_{r / 2}(u)$. Then

$$
\operatorname{Pr}\left[\operatorname{diam}\left(G_{n}(B)\right) \geq 2\left(K_{1}+1\right) \ln n\right]=O\left(n^{-3}\right)
$$

where $G_{n}(B)$ is the induced subgraph of $G_{n}$ in $B$.

## Proof

Given $\tau_{0}$ and $N$, we consider the following process which generates a sequence of graphs $H_{s}=\left(W_{s}, F_{s}\right), s=1,2, \ldots, N$. (The meanings of $N, \tau_{0}$ will become apparant soon).
Time step 1
To initialize the process, we start with $H_{1}$ consisting of $\tau_{0}$ isolated vertices $y_{1}, \ldots, y_{\tau_{0}}$. Time step $s \geq 1$ : We add vertex $y_{s+\tau_{0}}$. We then add $\frac{m}{8000(\alpha+1)^{2}}$ random edges incident with $y_{s+\tau_{0}}$ of the form $\left(y_{s+\tau_{0}}, w_{i}\right)$ for $i=1,2, \ldots, \frac{m}{8000(\alpha+1)^{2}}$. Here each $w_{i}$ is chosen uniformly from $W_{s}$.

The idea is to couple the construction of $G_{n}$ with the construction of $H_{N}$ for $N \sim$ $\operatorname{Bi}\left(n-T, A_{r} / 4\right)$ and $\tau_{0}=\operatorname{Bi}\left(T, A_{r} / 4\right)$ such that whp $H_{N}$ is a subgraph of $G_{n}$ with vertex set $V_{n}(B)$. We are then going to show that whp $\operatorname{diam}\left(H_{N}\right) \leq 2\left(K_{1}+1\right) \ln n$, and therefore $\operatorname{diam}\left(G_{n}(B)\right) \leq 2\left(K_{1}+1\right) \ln n$.
To do the coupling we use two counters, $t$ for the steps in $G_{n}$ and $s$ for the steps in $H_{N}$ :

- Given $G_{\tau_{0}}$, set $s=0$. Let $W_{0}=V_{T}(B)$. Notice that $\tau_{0}=\left|W_{0}\right| \sim \operatorname{Bi}\left(T, A_{r} / 4\right)$ and that $\tau_{0} \leq K_{1} \ln n \mathbf{w h p}$.
- For every $t>T$.
- If $x_{t} \notin B$, do nothing in $H_{s}$.
- If $x_{t} \in B$, set $s:=s+1$. Set $y_{s+\tau_{0}}=x_{t}$. As we want $H_{N}$ to be a subgraph of $G_{n}$ we must choose the neighbors of $y_{s+\tau_{0}}$ among the neighbors of $x_{t}$ in $G_{n}$. Let $A$ be the set of vertices chosen by $x_{t}$ in $V_{t}(B)$. Notice that $|A|$ stochastically dominates $a_{t} \sim \operatorname{Bi}\left(m, \frac{D_{t}(B)}{\max \left\{\alpha m A_{r} t, D_{t}\left(x_{t}\right)\right\}}\right)$. If $\frac{D_{t}(B)}{\max \left\{\alpha m A_{r} t, D_{t}\left(x_{t}\right)\right\}} \geq \frac{1}{50(\alpha+1)}$, then $a_{t}$ stochastically dominates $b_{t} \sim \operatorname{Bi}\left(m, \frac{1}{50 \alpha}\right)$ and so whp is at least $\frac{m}{100(\alpha+1)}$. If $\frac{D_{t}(B)}{\max \left\{\alpha m A_{r} t, D_{t}\left(x_{t}\right)\right\}}<\frac{1}{50(\alpha+1)}$ we declare failure, but as we see below this is unlikely to happen.
For any $R>0$,

$$
\begin{align*}
m\left|V_{t}\left(B_{R}(w)\right)\right| \leq D_{t}\left(B_{R}(w)\right)=m\left|V_{t}\left(B_{R}(w)\right)\right|+ & D_{t}^{-}\left(B_{R}(w)\right) \\
& \leq 2 m\left|V_{t}\left(B_{R+r}(w)\right)\right| \tag{18}
\end{align*}
$$

where $D_{t}^{-}\left(B_{R}(w)\right)$ is the sum over vertices $x \in B_{R}(w)$ of the of the in-degree $\operatorname{deg}_{t}(x)-m$ of $x$.
Now $\mid V_{t}\left(B_{R}(w) \mid \sim \operatorname{Bi}\left(t,(R / r)^{2} A_{r}\right)\right.$ and so

$$
\begin{align*}
\operatorname{Pr}\left(D_{t}\left(x_{t}\right) \geq 8 m A_{r} t \text { OR } D_{t}(B) \notin\right. & {\left[m A_{r} t / 5,3 m A_{r} t\right] } \\
& \left.\quad \text { OR }\left|V_{t}(B)\right|<A_{r} t / 5\right) \leq n^{-K_{1} / 100} \tag{19}
\end{align*}
$$

So we assume that $G_{t}$ is such that the event described in (19) does not happen. Thus each vertex of $B$ has probability at least $\frac{m}{8(\alpha+1) m A_{r} t} \geq \frac{1}{40(\alpha+1)\left|V_{t}(B)\right|}$ of being chosen under preferential attachment.
Thus, as insightfully observed by Bollobás and Riordan [11] we can legitimately start the addition of $x_{t}$ in $G_{t}$ by choosing $\frac{m}{8000(\alpha+1)^{2}}$ random neighbours uniformly in $B$.

Notice that $N$, the number of times $s$ is increased, is the number of steps for which $x_{t} \in B$, and so $N \sim \operatorname{Bi}\left(n-T, A_{r} / 4\right)$.
Now we are ready to show that $H_{N}$ is connected whp.
By Chernoff's bound we have that

$$
\operatorname{Pr}\left[\left|\tau_{0}-\frac{K_{1}}{4} \ln n\right| \geq \frac{K_{1}}{8} \ln n\right] \leq 2 n^{-K_{1} / 48}
$$

and

$$
\operatorname{Pr}\left[N \leq \frac{1}{3}(\ln n)^{2}\right] \leq e^{-c(\ln n)^{2}}
$$

for some $c>0$. Therefore, we can assume $\ln n \leq \tau_{0} \leq K_{1} \ln n$ and $N \geq \frac{1}{3}(\ln n)^{2}$. Let $X_{s}$ be the number of connected components of $H_{s}$. Then

$$
X_{s+1}=X_{s}-Y_{s}, \quad X_{0}=\tau_{0}
$$

where $Y_{s} \geq 0$ is the number of components (minus one) collapsed into one by $y_{s+\tau_{0}}$. So

$$
\operatorname{Pr}\left[Y_{s}=0 \mid H_{s}\right] \leq \sum_{i=1}^{X_{s}}\left(\frac{c_{i}}{s+\tau_{0}}\right)^{m / 8000(\alpha+1)^{2}}
$$

where the $c_{i}$ are the component sizes of $H_{s}$. If $s<2 K_{1} \ln n$ then because $m \geq K \ln n$, we have

$$
\operatorname{Pr}\left[Y_{s}=0 \mid X_{s} \geq 2\right] \leq 2\left(1-\frac{1}{s+\tau_{0}}\right)^{m / 8000(\alpha+1)^{2}} \leq 2 e^{-m /\left(8000(\alpha+1)^{2}\left(s+\tau_{0}\right)\right)} \leq 1 / 10
$$

So $X_{s}$ is stochastically dominated by the random variable $\max \left\{1, \tau_{0}-Z_{s}\right\}$ where $Z_{s} \sim$ $\operatorname{Bi}(s, 9 / 10)$. We then have

$$
\operatorname{Pr}\left[X_{2 K_{1} \ln n}>1\right] \leq \operatorname{Pr}\left[Z_{2 K_{1} \ln n}<\tau_{0}\right] \leq \operatorname{Pr}\left[Z_{2 K_{1} \ln n}<K_{1} \ln n\right] \leq n^{-3}
$$

And therefore

$$
\operatorname{Pr}\left[H_{2 K_{1} \ln n} \text { is not connected }\right] \leq n^{-3}
$$

Now, to obtain an upper bound on the diameter, we run the process of construction of $H_{N}$ by rounds. The first round consists of $2 K_{1} \ln n$ steps and in each new round we double the size of the graph, i.e. it consists of as many steps as the total number of steps of all the previous rounds. Notice that we have less than $\ln n$ rounds in total. Let $\mathcal{A}$ be the event for all $i>0$ every vertex created in the $(i+1)^{t h}$ round is adjacent to a vertex in $H_{2^{i-1} K_{1} \ln n}$, the graph at the end of the $i^{t h}$ round.
On the event $\mathcal{A}$, every vertex in $H_{N}$ is at distance at most $\ln n$ of $H_{2 K_{1} \ln n}$ whose diameter is not greater than $2 K_{1} \ln n$. Thus, the diameter of $H_{N}$ is smaller than $2\left(K_{1}+1\right) \ln n$. Now, we have that if $v$ is created in the $(i+1)^{t h}$ round,

$$
\operatorname{Pr}\left[v \text { is not adjacent to } H_{2^{i-1} K_{1} \ln n}\right] \leq\left(\frac{1}{2}\right)^{m}
$$

Therefore

$$
\operatorname{Pr}[\neg \mathcal{A}] \leq\left(\frac{1}{2}\right)^{m} n(\ln n) \leq \frac{\ln n}{n^{K \ln 2-1}}
$$

To finish the proof of connectivity and the diameter, let $u, v$ be two vertices of $G_{n}$. Let $C_{1}, C_{2}, \ldots, C_{M}, M=O(1 / r)$ be a sequence of spherical caps of radius $r / 4$ such that $u$ is the center of $C_{1}, v$ is the center of $C_{M}$ and such that the centers of $C_{i}, C_{i+1}$ are distance $\leq r / 2$ apart. The intersections of $C_{i}, C_{i+1}$ have area at least $A_{r} / 40$ and so whp each
intersection contains a vertex. Using Lemma 6 we deduce that whp there is a path from $u$ to $v$ in $G_{n}$ of size at most $O(\ln n / r)$.
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